A Short Introduction To Likelihood in Phylogenetics

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Charles Darwin: Evolutionary Relationships





Charles Darwin (1809-1882)









I.

| | А | G | С | С | А | T | G | С | А | G | |
|------|---|---|--------|--------|--------|--------|---|---|---|--------|--|
| | A | G | с | С | A | * C | G | С | A | G | |
| | А | G | с | С | ♦ G | с | G | С | A | G | |
| | A | G | С | С | ť | ç | G | С | A | G | |
| | A | G | С | С | с | * G | G | С | А | G | |
| time | A | G | ç | * A | с | G | G | С | А | G | |
| | A | G | * G | A | с | G | G | С | А | G | |
| | A | G | G | A | с | Å | G | С | А | G | |
| | A | G | Ģ | A | с | A | G | С | А | * A | |
| | A | G | č | A | с | A | G | С | А | A | |
| , | A | G | с | А | с | Ť | G | с | А | A | |

Some Notation



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Note: branch = edge = split, external node = leaf = taxon are used interchangebly.

| Data | Method | Evaluation Criterion | | |
|---------------------------|---|-------------------------|--|--|
| (| Maximum Parsimony | Parsimony | | |
| Characters (Alignment) | Statistical Approaches: Likelihood, Bayesian | Evolutionary Models | | |
| Distances | Distance Methods |) | | |

The most simple tree could be seen as two sequences and the distance between them.

Distances can be computed in various ways...

The most simple tree: How to get distances?

The most simple tree could be seen as two sequences and the distance between them.

Distances can be computed in various ways...

Usually via Maximum Likelihood.



Ronald Fisher (1890-1962)



Joe Felsenstein (born 1942)

Given a box with 3 coins with different levels of fairness $(\frac{1}{3}, \frac{1}{2}, \frac{2}{3} \text{ heads})$

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We take out one coin and toss 20 times:

H, T, T, H, H, T, T, T, T, H, T, T, H, T, H, T, T, H, T, T

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Probability

 $p(k \text{ heads in } n \text{ tosses}|\theta)$

Aim: The ML approach seaches for that parameter set θ for the generating process which maximizes the probability of our given data.

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Hence, "likelihood flips the probability around."

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ProbabilityLikelihood $p(k \text{ heads in } n \text{ tosses}|\theta) \equiv L(\theta|k \text{ heads in } n \text{ tosses})$ $= \binom{n}{k} \theta^k (1-\theta)^{n-k}$
(here binomial distribution)

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Three coin case

$$L(\theta|7 \text{ heads in } 20) = {\binom{20}{7}}\theta^7 (1-\theta)^{13}$$

for each coin $\theta \in \{\frac{1}{3}, \frac{1}{2}, \frac{2}{3}\}$



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For infinitely many coins $\theta = (0...1)$



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ML estimate: $L(\hat{\theta}) = 0.1844$ where coin shows $\hat{\theta} = 0.35$ heads

While the coin tossing example might look easy, in phylogenetic analysis, the parameter (set) θ comprises:

- evolutionary model
- its parameters
- tree topology
- its branch lengths

That means a high dimensional optimization problem. Hence, some parameters are often estimated/set separately. • Evolution is usually modeled as a

stationary, time-reversible Markov process.

• What does that mean?

Markov Process

The (evolutionary) process evolves without memory, i.e. sequence S_2 mutates to S_3 during time t_{n+1} independent of state of S_1 .



Stationary:

The overall character frequencies π_j of the nucleotides or amino acids are in an equilibrium and remain constant.

Time-Reversible:

Mutations in either direction are equally likely

$$\pi_i \cdot P_{ij}(t) = P_{ji}(t) \cdot \pi_j$$

This means a mutation is as likely as its back mutation.

$$P(i \to j) = P(i \leftarrow j) \tag{JC69}$$

Substitution Models

Evolutionary models are often described using a substitution rate matrix R and character frequencies Π . Here, 4×4 matrix for DNA models:



$$R = \begin{pmatrix} A & C & G & T \\ - & a & b & c \\ a & - & d & e \\ b & d & - & f \\ c & e & f & - \end{pmatrix}$$

$$\mathbf{\Pi} = (\pi_A, \pi_C, \pi_G, \pi_T)$$

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From R and Π we reconstruct a substitution probability matrix P, where $P_{ij}(t)$ is the probability of changing $i \rightarrow j$ in time t.





rate heterogeneity: invariant sites, Γ -distributed rates, mixed.

Generally this is the same for protein sequences, but with 20×20 matrices. Some protein models are:

- Poisson model ("JC69" for proteins, rarely used)
- Dayhoff (Dayhoff et al., 1978, general matrix)
- JTT (Jones et al., 1992, general matrix)
- WAG (Whelan & Goldman, 2000, more distant sequences)
- VT (Müller & Vingron, 2000, distant sequences)
- mtREV (Adachi & Hasegawa, 1996, mitochondrial sequences)
- cpREV (Adachi et al., 2000, cloroplast sequences)
- mtMAM (Yang et al., 1998, Mammalian mitochondria)
- mtART (Abascal et al., 2007, Arthropod mitochondria)
- rtREV (Dimmic et al., 2002, reverse transcriptases)
- . . .
- BLOSUM 62 (Henikoff & Henikoff, 1992) \rightarrow for database searching

Computing ML Distances Using $P_{ij}(t)$

The Likelihood of sequence s evolving to s' in time t:

$$L(t|s \rightarrow s') = \prod_{i=1}^{m} \left(\Pi(s_i) \cdot P_{s_i s'_i}(t) \right)$$

Likelihood surface for two sequences under JC69:

GATCCTGAGAGAAATAAAC = s'GGTCCTGACAGAAATAAAC = s

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Note: we do not compute the probability of the distance t but that of the data $D = \{s, s'\}$.



Jukes-Cantor Correction for Multiple Mutations



time










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Given a tree with branch lengths and sequences for all nodes, the computation of likelihood values for trees is straight forward. Unfortunately, we usually have no sequences for the inner nodes (ancestral sequences).

Hence we have to evaluate every possible labeling at the inner nodes:

$$L \begin{pmatrix} c \\ c \end{pmatrix} = L \begin{pmatrix} c \\ c \end{pmatrix} + L \begin{pmatrix} c \\ c \end{pmatrix} + \cdots + L \begin{pmatrix} c \\ c \end{pmatrix} + L \begin{pmatrix} c \\ c$$

for every column in the alignment...but there is a faster algorithm. (Peeling Algorithm by Felsenstein, 1981)

For a single alignment column



For a single alignment column and a given tree:



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with all
$$d_x = 0.1$$
 and $P_{ij}(0.1) = \begin{cases} .9068 & i = j \\ .0313 & i \neq j \end{cases}$ (JC)

For a single alignment column and a given tree:

Likelihoods of nucleotides i at inner nodes:

$$L_{5}(i) = [P_{iC}(d_{1}) \cdot L_{1}(C)] \cdot [P_{iG}(d_{2}) \cdot L_{2}(G)]$$



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Likelihoods of nucleotides *i* at inner nodes:

 $L_{5}(i) = [P_{iC}(d_{1}) \cdot L_{1}(C)] \cdot [P_{iG}(d_{2}) \cdot L_{2}(G)]$

$$L_{6}(i) = \prod_{\nu = \{3,4,5\}} \left[\sum_{j = \{ACGT\}} P_{ij}(d_{\nu}) \cdot L_{\nu}(j) \right]$$

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Site-Likelihood of an alignment column k:

$$L^{(k)} = \sum_{i \in \{ACGT\}} \pi_i \cdot L_6(i) = 0.005489$$

with all $d_x = 0.1$ and $P_{ij}(0.1) = \begin{cases} .9068 & i = j \\ .0313 & i \neq j \end{cases}$ (JC)



Likelihoods of Trees (multiple columns)



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Considering this tree with n = 4 sequences of length m = 3 the tree likelihood of this tree CCG GGC is 0.1 0.1 CCC CCC $\mathcal{L}(T) = \prod L^{(k)}$ $= 0.005489^2 \cdot 0.005489$ k=10.1 0.0000001653381 0.

Likelihoods of Trees (multiple columns)

Considering this tree with n = 4 sequences of GGC length m = 3 the tree likelihood of this tree CCG is 0.1 0.1 $\mathcal{L}(T) = \prod_{k=1}^{m} L^{(k)} = 0.005489^2 \cdot 0.005489$ CCC CCC k=10.1 0.0000001653381 0. or the log-likelihood $\ln \mathcal{L}(T) = \sum_{k=1}^{m} \ln L^{(k)} = -15.61527$









(1.) Choose a branch. (2.) Move the virtual root to an adjacent node.(3.) Compute all partial likelihoods recursively. (4.) Adjust the branch length to maximize the likelihood value.



Repeat this for every branch until no better likelihood is gained.

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This is based on the Pulley-Principle (Felsenstein, 1981) which states that the root can be moved on the tree but the likelihood doesn't change.





Number of Trees to Examine...



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$$B(n) = \frac{(2n-5)!}{2^{n-3}(n-3)!}$$

$$B(10) = 2027025$$

$$B(55) = 2.98 \cdot 10^{84}$$

$$B(100) = 1.70 \cdot 10^{182}$$



Exhaustive Search: guarantees to find the optimal tree, because all trees are evaluated, but not feasible for more than 10-12 taxa.
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but often not for current-day datasets.

Heuristics: cannot guarantee to find the optimal tree, but are at least able to analyze large datasets.













Is also used for other (non-ML) methods like parsimony.

What if we have multiple maxima in the likelihood surface?



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Tree rearrangements to escape local maxima.





From a current tree construct other trees by rearranging its subtrees and evaluates all resulting trees. Repeat with the best tree found, until no better tree can be found. This also used for other (non-ML) methods, like parsimony.



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- Usually programs deliver a single (best) tree, but without confidence values for the subtrees.
- How can we assess reliability for the subtree?

- Bootstrapping creates many pseudo-alignments by sampling alignment columns with replacement from the original alignment.
- From the pseudo-alignment we reconstruct trees.
- From the trees we collect and count all splits.
- From the splits we construct a consensus tree.
- **Definition:** A split *A*|*B* in the tree is the bipartition of the leaves/taxa into two subsets *A* and *B* induced by removing an edge or branch from the tree.
- Definition: Two splits A|B and C|D are compatible, i.e. not contradictory, if at least one intersection of A ∩ C, A ∩ D, B ∩ C, B ∩ D is empty.













- Strict consensus: contains all splits occuring in all input tree.
- Semi-strict consensus: contains all splits which are not contradicted by any tree.
- Majority consensus M_{ℓ} : contains all splits which occur in more than ℓ input trees, where $\ell \geq 50\%$ typically exactly 50%.



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The Bayesian approach asks the right question in a hypothesis testing procedure, namely, "What is the probability that this hypothesis is true, given the data?" rather than the classical approach, which asks a question like, "Assuming that this hypothesis is true, what is the probability of the observed data?"

Ewens & Grant: Statistical Methods in Bioinformatics (2010)

Bayesian statistics tries to determine the probability of the hypothesis:

$$\Pr(hypothesis|data) = rac{\Pr(hypothesis) imes \Pr(data|hypothesis)}{\Pr(data)}$$

where

- Pr(*data*|*hypothesis*) is the 'likelihood'
- Pr(*hypothesis*) is the prior distribution
- Pr(*hypothesis*|*data*) is the posterior distribution
- Pr(*data*) is the marginal likelihood

Bayesian Phylogenetics

Bayesian statistics tries to determine the probability of the hypothesis (i.e. tree topologies, branch lengths, evolutionary model with its parameters):

$$p(T, M|D) = \frac{p(T) \times p(M) \times \Pr(D|T, M)}{\Pr(D)}$$

where

- Pr(D|T, M) is the 'likelihood'
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- Pr(*T*, *M*|*data*) is the posterior distribution
- Pr(D) is the marginal likelihood of the data
- The output of a Bayesian evolutionary analysis is a probability distribution on trees and parameter values.